

Conceptualizing and testing moderated effects

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Epi 222: Health Disparities Research Methods
May 14, 2009

What is moderation?

- . When the relationship between an X variable and a Y variable changes as a function of a third variable (i.e., another X variable).
- . When the effect of one explanatory variable depends on the level of another explanatory variable.
- . When two (or more) variables considered in combination have a joint effect on the outcome that is greater than the sum of their individual effects.
- . A statistical interaction.

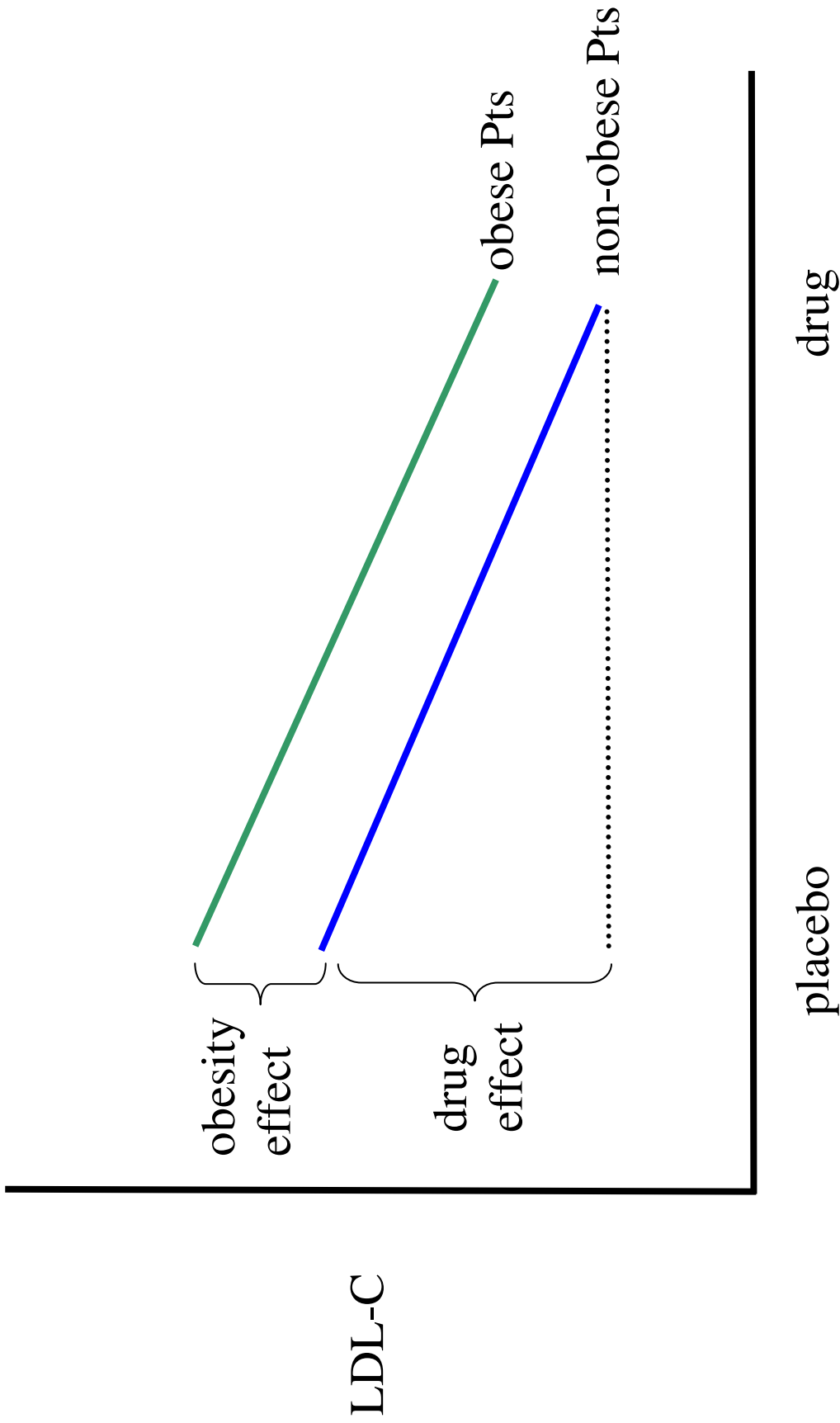
Why bother with moderation?

Most empirical models of health disparities focus exclusively on main effects

If an interaction effect exists, then a main effects model is misspecified,
leading to biased effect estimates

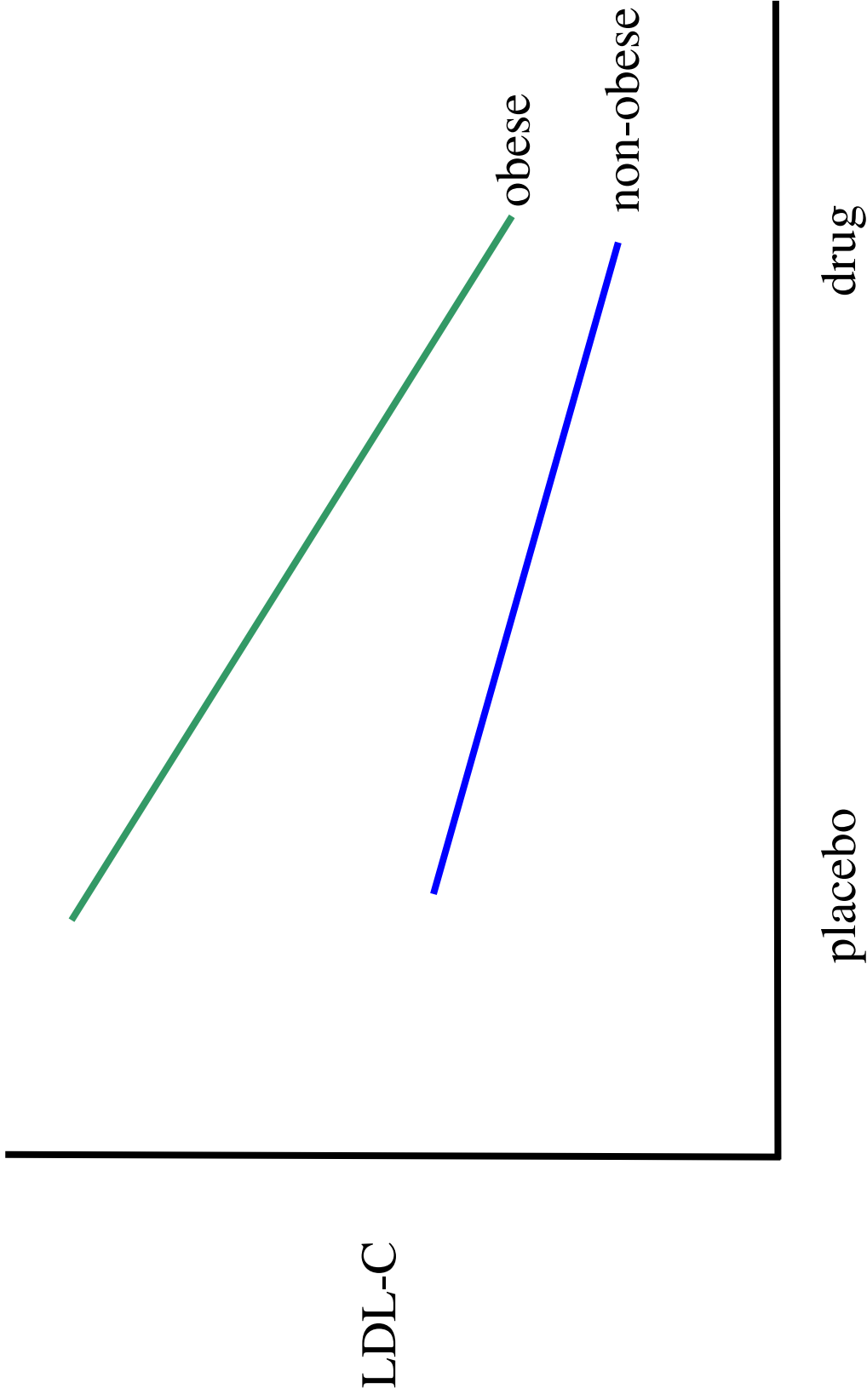
Little empirical progress has been made toward explaining health disparities.
A focus on potential interaction effects may help to ameliorate...

Graphical examples: main effects only, no interaction



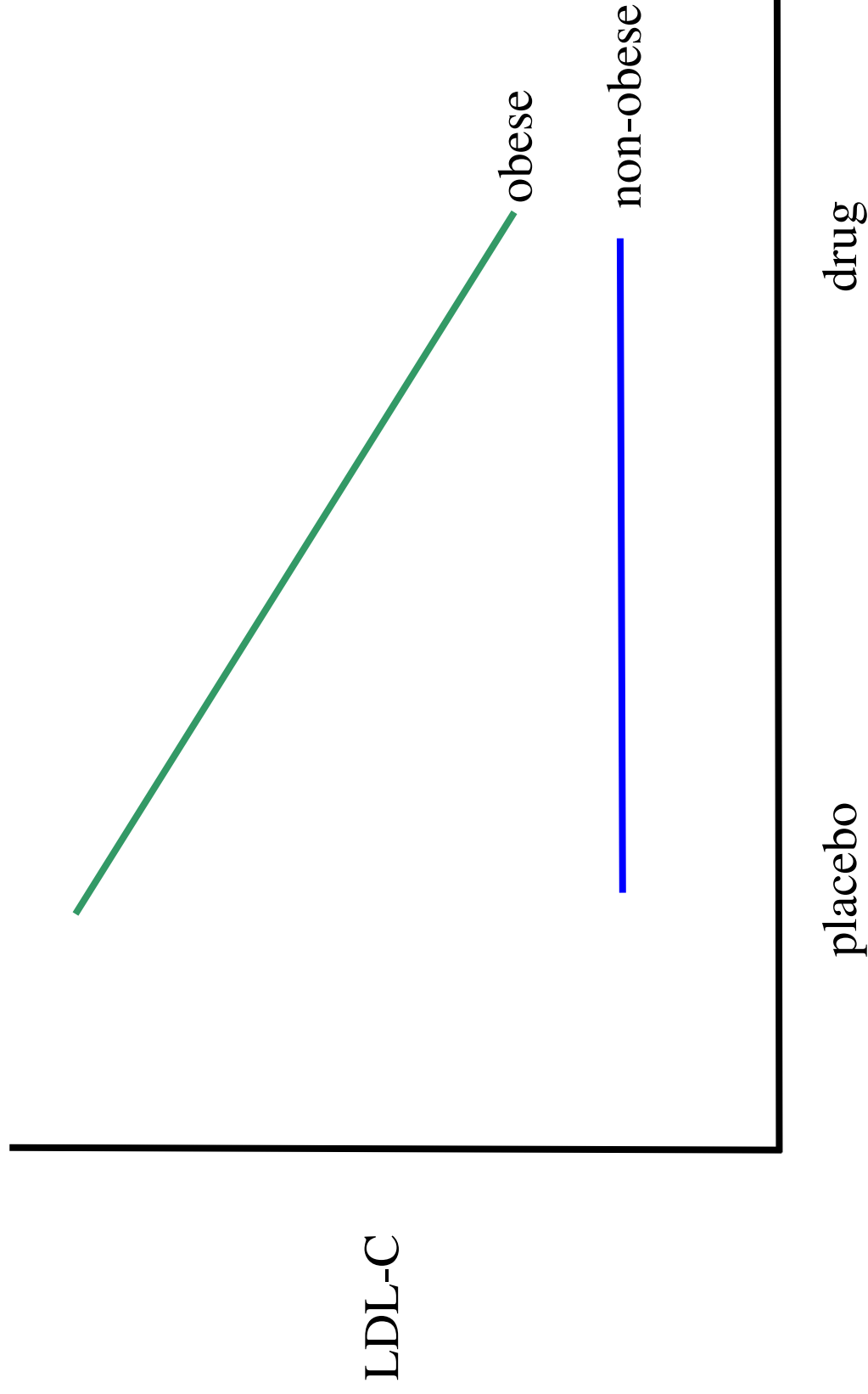
. What can be said about the effect of the drug? the effect of obesity?

Graphical examples: interaction

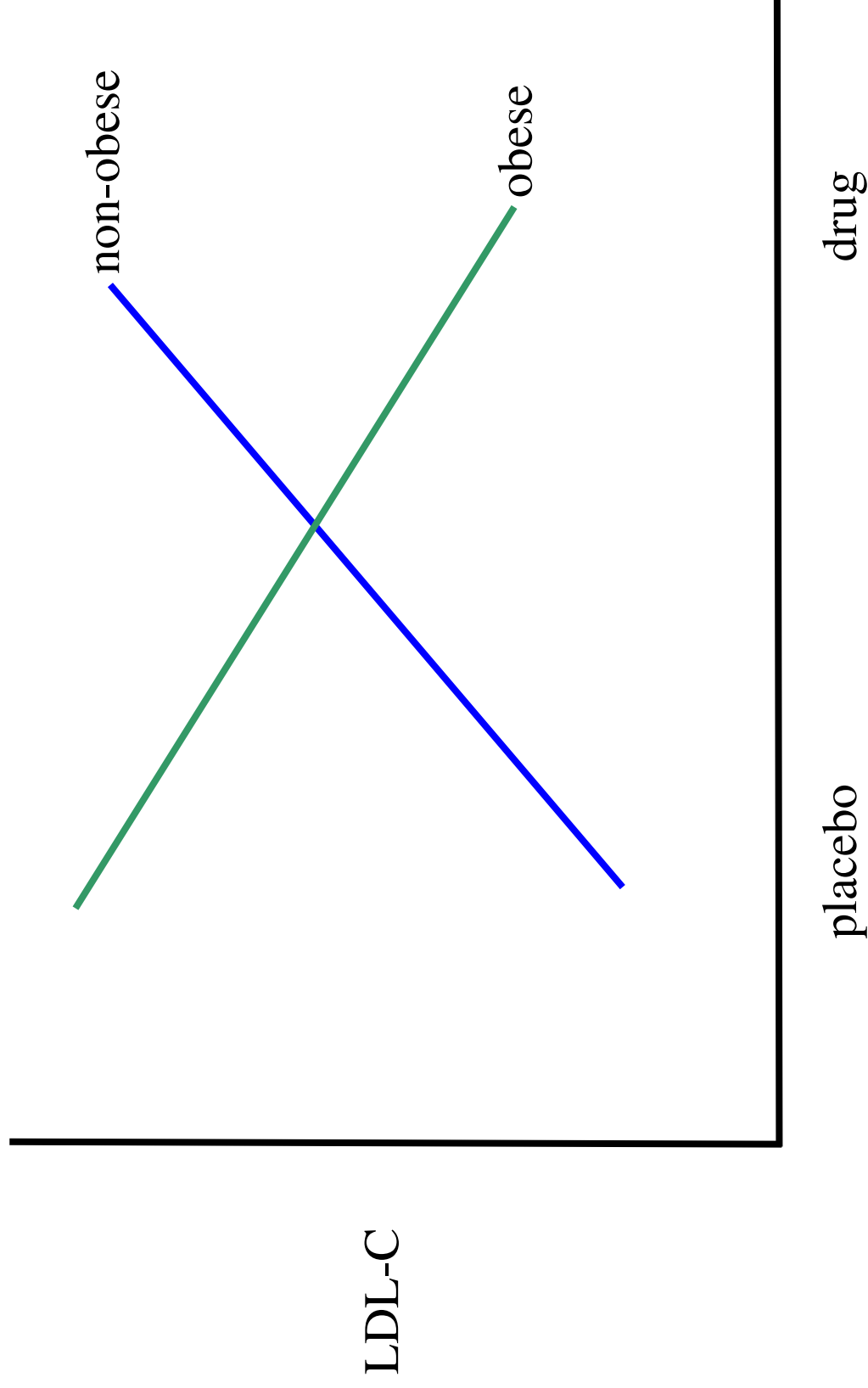


. What can be said about the effect of the drug? the effect of obesity?

Graphical examples: other possibilities



Graphical examples: other possibilities



Graphical examples: other possibilities

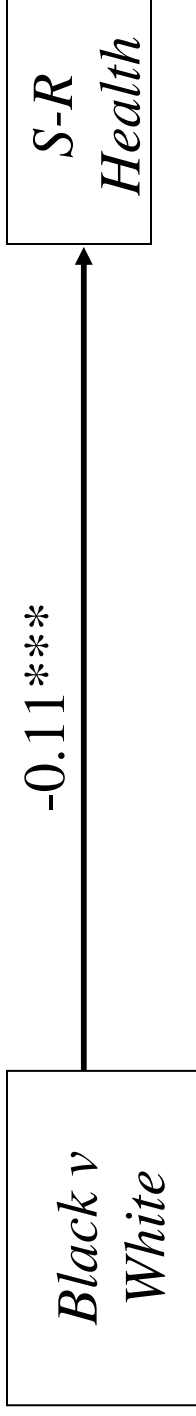
Summary

Interactions can take many forms, but the shared characteristic is that the association between X and Y is not constant

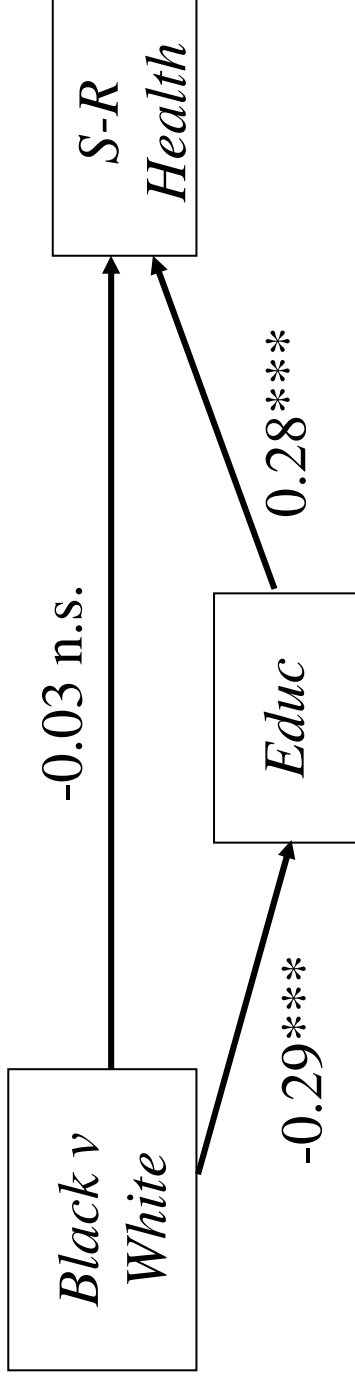
- the magnitude of that association depends on a the value of an additional variable (or set of variables).
- the magnitude of the association between X and Y significantly differs as a function of the additional variable(s)

Introduction: Mediation models are main effects models

Consider the following total effect model

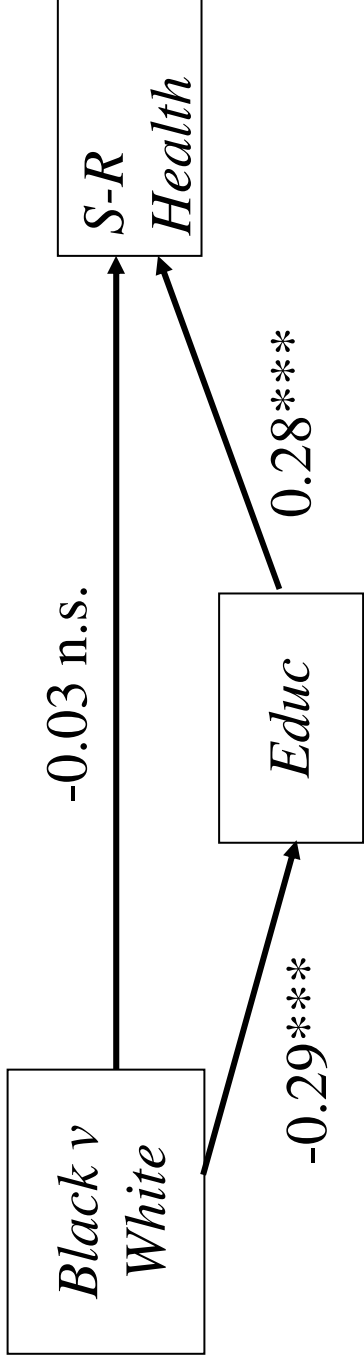


And one possible corresponding mediation model



Introduction: Mediation models are main effects models

- The mediation model...



...assumes that the effect of Race is constant at all education levels

- Defending the estimated conditional effect of Race rests upon this assumption
- This assumption can be tested by estimating and testing interaction effects.

My goal is to provide a conceptual introduction to testing interactions.

Example data: EPESE

Established Populations for Epidemiologic Studies of the Elderly (EPESE)

- Duke site
- Probability sample
- Baseline data collected in 1982
- 65 years and older
- 54% African American
- $N \approx 2900$ (with complete data on key variables)

Example data: EPESE

Outcome: self-rated health

*Compared to other people your own age,
would you say that your general health is
excellent, good, fair, poor?*

How would you rate your health at the present time?

code	label	frequency	
		4-category	binary
1	poor	570	1855
2	fair	1285	
3	good	1546	2086
4	excellent	540	

For demonstration,
self-rated health can be treated as continuous, categorical, or binary.

Example data: EPESE

Explanatory variables

Race

code	label	frequency
0	White	1820
1	Black	2121

Education

code	label	frequency	
		4-category	binary
0	<8	1715	3030
1	8-11	1315	
2	= HS	335	911
3	> HS	576	

For demonstration,
education can be treated as continuous, categorical, or binary.

Types of moderation models covered

- Example 1: two binary X variables with continuous Y : pooled data
- Example 1A: two binary X variables with continuous Y : stratified analyses
- Example 2: two binary X variables with a binary Y : pooled data
- Example 3: a binary & a continuous X with a continuous Y
- Example 4: A binary & a categorical X with a continuous Y

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Example 1: Two binary X variables with a continuous Y

Self-rated health means & mean differences as a function of two binary variables

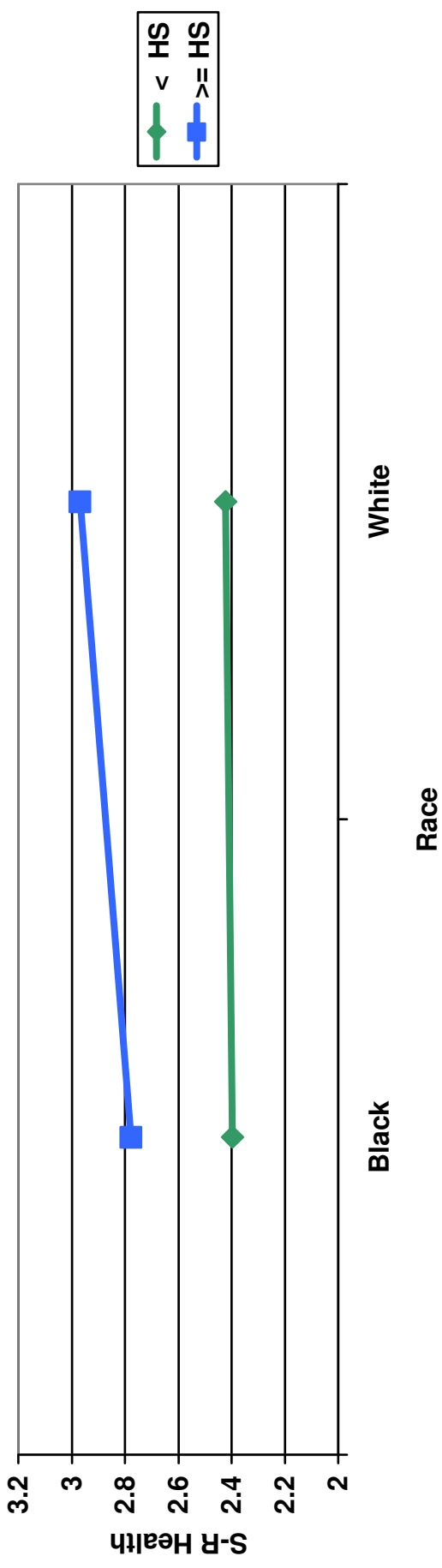
race	education		Δ
	<HS (code=0)	\geq HS (code=1)	
White (code=0)	2.422	2.969	0.547
Black (code=1)	2.396	2.777	0.381
Δ	-0.026	-0.192	-0.166

- . The simple differences (pink and gray) represent main effects.
 - . The difference between the differences (yellow) is the interaction effect.
- Does the difference between <HS and \geq HS differ across the races?*
- Does the difference between the races differ across the education level?*

Example 1: Two binary X variables with a continuous Y

Self-rated health means & mean differences as a function of two binary variables

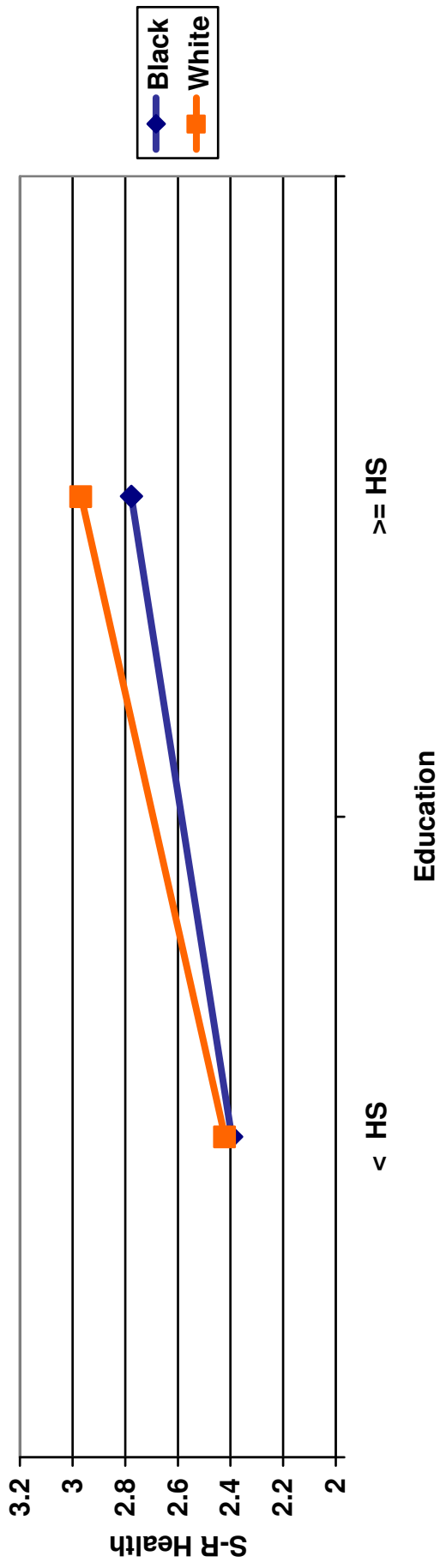
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Example 1: Two binary X variables with a continuous Y

$$\text{S-R}_Hx = \mathbf{B}_0 + \mathbf{B}_1 \times \text{Race} + \mathbf{B}_2 \times \text{Educ} + \mathbf{B}_3 \times \text{Race} \times \text{Educ} + \text{resid.}$$

Parameter	DF	Estimate	StdErr	t	p
B0 (int)	1	2.4218	0.0252	96.18	<.0001
B1 (race)	1	-0.0259	0.0325	-0.80	0.4252
B2 (educ)	1	0.5471	0.0435	12.59	<.0001
B3 (rxeduc)	1	-0.1664	0.0697	-2.39	0.0171
<i>(custom test)</i>					
educ Blacks	1	0.3807	0.0546	6.99	<.0001

Self-rated health means & mean differences as a function of two binary variables

race	education		Δ
	<HS (code=0)	\geq HS (code=1)	
White (code=0)	2.4218	2.9689	0.5471
Black (code=1)	2.3959	2.7766	0.3807
Δ	-0.0259	-0.1923	-0.1664

Example 1: Two binary X variables with a continuous Y

Summary

- . among Black respondents, those with \geq HS averaged 0.381 points higher on the self-rated health outcome compared to those with $<$ HS, $p < .0001$.
- . among White respondents, those with \geq HS averaged 0.547 points higher on the self-rated health outcome compared to those with $<$ HS, $p < .0001$.
- . A significant interaction existed between race and education: the effect of education was significantly stronger for Whites than for Blacks, $p < .02$.

The simple approach

If the interaction is significant, then

- report the p-value for the interaction and
- report on the effect of education within each race, or
- report on the effect of race within each education level.

Be cautious when reporting main effects.

Make sure you are very clear about them.

Types of moderation models covered

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Example 1A: Two binary X variables with a continuous Y

Stratified analyses (fit the following model within each race-specific sample)

$$\mathbf{S-R_Hx} = \mathbf{B_0} + \mathbf{B_2 \times Educ} + \mathbf{resid.}$$

Whites (n=1820)

Parameter	DF	Estimate	StdErr	t	p
$B0_W$ (int)	1	2.4218	0.0252	96.04	<.0001
$B2_W$ (educ)	1	0.5471	0.0435	12.57	<.0001

Blacks (n=2121)

Parameter	DF	Estimate	StdErr	t	p
$B0_B$ (int)	1	2.3959	0.0205	116.93	<.0001
$B2_B$ (educ)	1	0.3807	0.0545	6.99	<.0001

t-test of interaction effect

$$\begin{aligned} t &= (B2_B - B2_W) / \text{SQRT}(SE_B^2 + SE_W^2) \\ &= (0.3807 - 0.5471) / \text{SQRT}(.0545^2 + .0435^2) \\ &= -2.39 \end{aligned}$$

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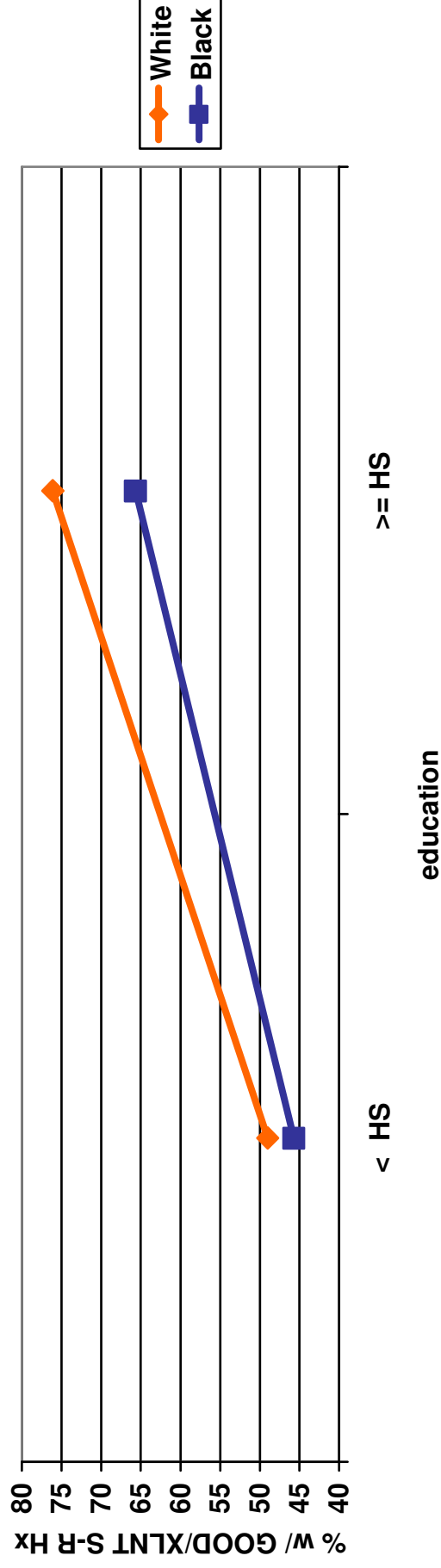
Example 3: a binary & a continuous X with a continuous Y

Example 4: A binary & a categorical X with a continuous Y

Example 2: Two binary X variables with a binary Y

% with good/excellent self-rated health as a function of two binary variables

race	education		OR
	<HS	≥HS	
White	48.97%	76.10%	3.319
Black	45.69%	65.67%	2.274
OR	0.877	0.601	0.685 (OR ratio)



Example 2: Two binary X variables with a binary Y

% with good/excellent self-rated health as a function of two binary variables

race	education		Δ logit
	<HS	≥HS	
White	48.97% (-0.0414)	76.10% (1.1584)	1.1998
Black	45.69% (-0.1729)	65.67% (0.6484)	0.8213
Δ logit	-0.1315	-0.5100	-0.3785

logit?

logit = $\ln(\pi/(1-\pi))$, where π is the response probability

E.g., the logit of .4897 = $\ln(0.4897 \div (1 - 0.4897)) = -0.0414$

Example 2: Two binary X variables with a binary Y

% with good/excellent self-rated health as a function of two binary variables

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	<HS	≥HS	
White	48.97% (-0.0414)	76.10% (1.1584)	1.1998
Black	45.69% (-0.1729)	65.67% (0.6484)	0.8213
Δ logit	-0.1315	-5100	-0.3785

Parameter	DF	Estimate	Error	t	p
B0	1	-0.0414	0.0575	0.72	0.4722
B1 (race)	1	-0.1315	0.0743	1.77	0.0768
B2 (educ)	1	1.1998	0.1109	10.81	<.0001
B3 (race*educ)	1	-0.3785	0.1712	2.21	0.0271

OR[educ for Whites] = $\exp(1.1998) = 3.3194$, $p < .0001$

OR[educ for Blacks] = $\exp(0.8213) = 2.2735$, $p < .0001$

Example 2: Two binary X variables with a binary Y

Summary

- . among Black respondents, those with \geq HS had 2.27 higher odds of good/excellent self-rated health compared to those with $<$ HS, $p < .0001$.
- . among White respondents, those with \geq HS had 3.32 higher odds of good/excellent self-rated health compared to those with $<$ HS, $p < .0001$.
- . A significant interaction existed between race and education: the effect of education was significantly stronger for Whites than for Blacks, $p < .03$.

The simple approach

If the interaction is significant, then

- report the p-value for the interaction effect, and
- report on the effect (OR) of education within each race, or
- report on the effect (OR) of race within each education level.

Be cautious when reporting main effects.

Make sure you are very clear about them.

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Example 3: A binary & continuous X with a continuous Y

Mean self-rated health

Whites: 2.605

Blacks: 2.450

Education

Treated as a continuous variable with values 0, 1, 2, and 3

mean = 0.942

std dev = 1.050

Example 3: A binary & continuous X with a continuous Y

$$\text{S-R}_Hx = \mathbf{B}_0 + \mathbf{B}_1 \times \text{Race} + \mathbf{B}_2 \times \text{Educ} + \mathbf{B}_3 \times \text{Race} \times \text{Educ} + \text{resid.}$$

Parameter	DF	Estimate	StdErr	t	p
B0 (int)	1	2.2863	0.0315	72.64	<.0001
B1 (race)	1	0.0509	0.0391	1.30	0.1928
B2 (educ)	1	0.2530	0.0190	13.33	<.0001
B3 (rxex)	1	-0.0845	0.0276	3.06	0.0022

(custom test)

$$\text{educ (Black)} \quad 1 \quad \mathbf{0.1685} \quad 0.0200 \quad \cdot \quad 8.41 \quad <.0001$$

$$(\mathbf{0.1685} = \mathbf{0.2530} + \mathbf{-0.0845})$$

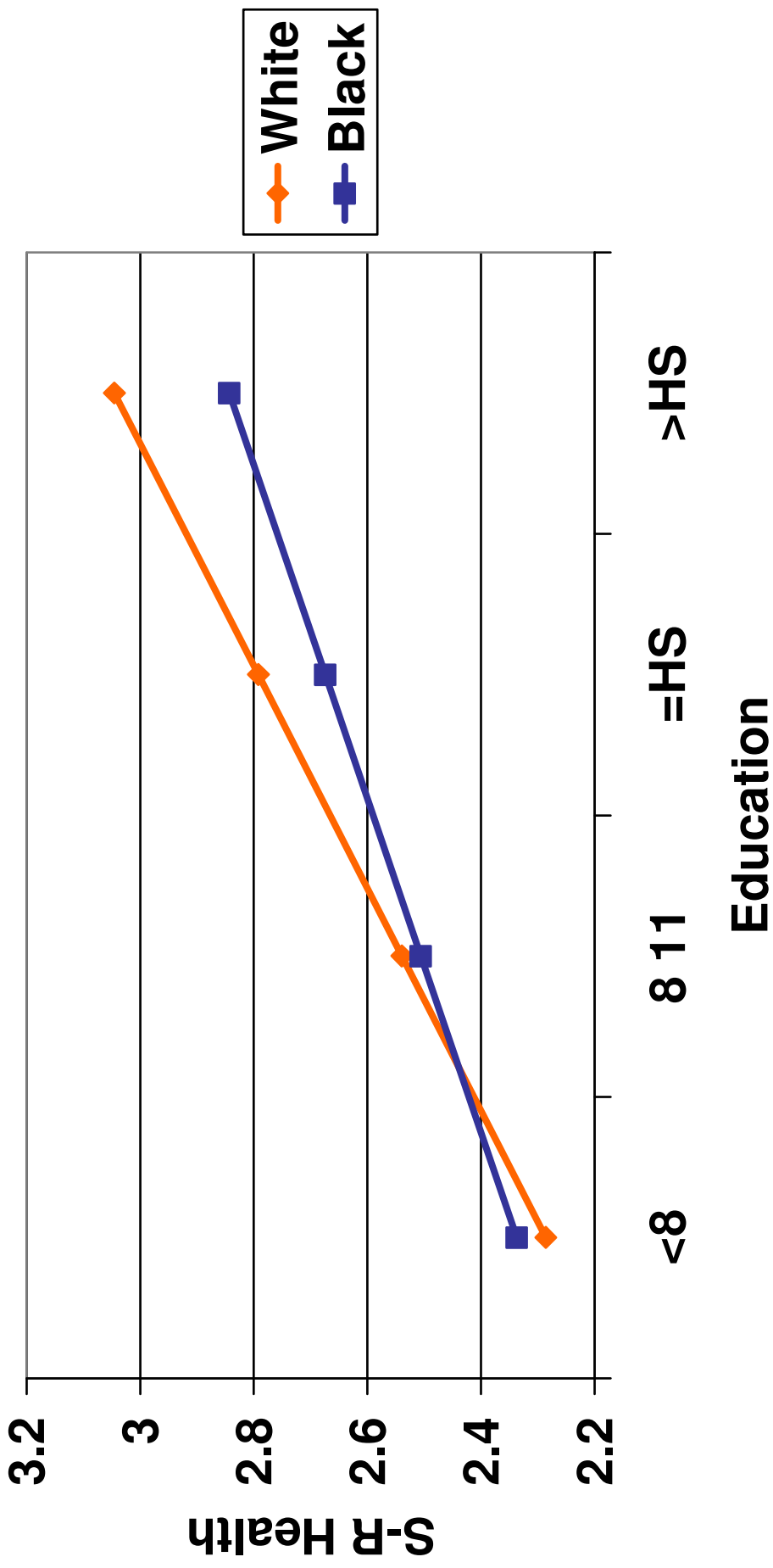
Model-predicted values of self-rated health

	education level		
	<8 (code=0)	8-11 (code=1)	=HS (code=2)
White (code=0)	2.2863	2.5393	2.7923
Black (code=1)	2.3372	2.5057	2.6742
Δ	0.0509	-0.0336	-0.1181
			>HS (code=3)
			3.0453
			2.8427
			-0.2026

0.2530 = **2.5393** – **2.2863**: effect of a one-unit increase in education for Whites

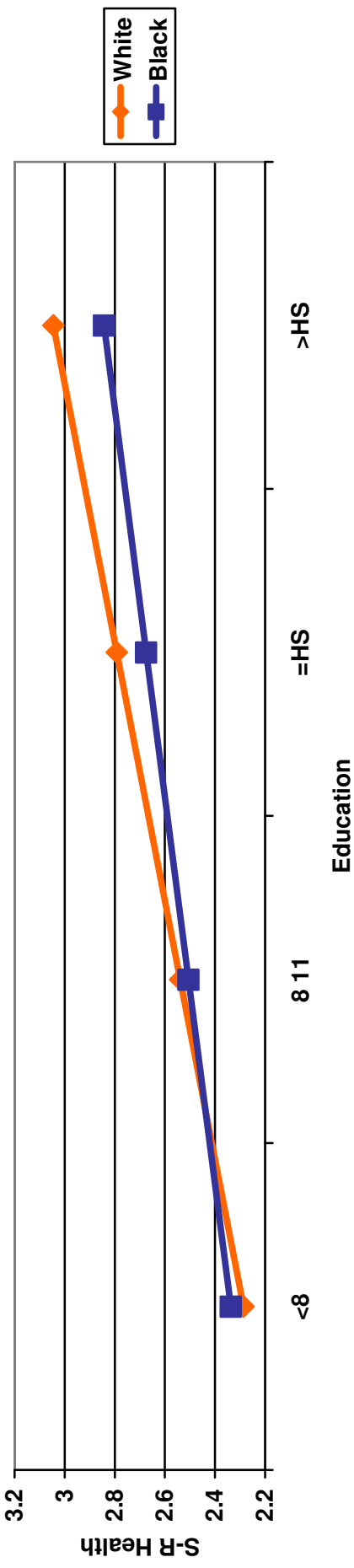
0.1685 = **2.5057** – **2.3372**: effect of a one-unit increase in education for Blacks

Example 3: A binary & continuous X with a continuous Y



What can be said about the effect of race? the effect of education?

Example 3: A binary & continuous X with a continuous Y



Summary

- . among White respondents, for every one-category increase in education the expected value of self-rated health increased by **0.2530** points, $p < .0001$.
- . among Black respondents, for every one-category increase in education the average self-rated health increased by **0.1685** points, $p < .0001$.
- . A significant interaction existed between race and education: the effect of education was significantly stronger for Whites than for Blacks, $p < .02$.

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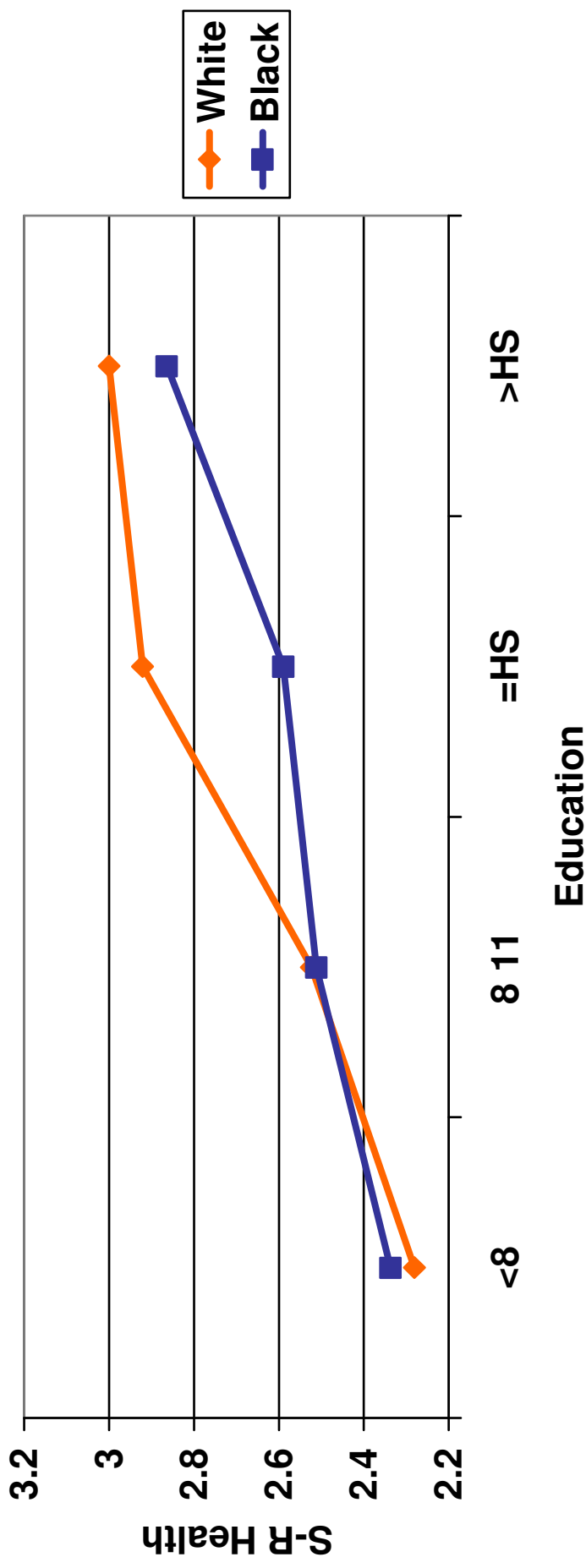
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Example 4: A binary & categorical X with a continuous Y

Observed mean values of self-rated health

	education level		
	<8	8-11	=HS
White (code=0)	2.2806	2.5235	2.9208
Black (code=1)	2.3375	2.5114	2.5895
			>HS
			3.0000
			2.8634



Example 4: A binary & categorical X with a continuous Y

This example has a binary race indicator and a 4-category education indicator.

Therefore, the effects of education and the race-by-education interaction will each have 3 degrees of freedom.

In such cases, I first look at the omnibus test of each effect.

Source	DF	χ^2	p
educ	3	180.28	<.0001
race*educ	3	13.80	0.0032

The interaction is significant

To describe the nature of the interaction, the choices are to

- (a) report race differences within each level of education, or
- (b) report education differences within each race

Example 4: A binary & categorical X with a continuous Y

Observed mean values of self-rated health

	education level		
	<8	8-11	=HS
White (code=0)	2.2806	2.5235	2.9208
Black (code=1)	2.3375	2.5114	2.5895
$\Delta(p)$	-0.0568 , $p=.217$.0120 , $p=.803$.3314 , $p=.001$

(*custom tests*)

Label	Estimate	Error	t	p
WvB: <8	-0.0568	0.0461	1.23	0.2174
WvB: 8-11	0.0120	0.0481	0.24	0.8025
WvB: =HS	0.3314	0.1055	3.14	0.0017
WvB: >HS	0.1366	0.0757	1.80	0.0713
W: HS v 8-11	0.3974	0.0651	6.10	<.0001
W: >HS v HS	0.0792	0.0721	1.10	0.2722
B: HS v 8-11	0.0780	0.0960	0.81	0.4162
B: >HS v HS	0.2739	0.1080	2.54	0.0112

Example 4: A binary & categorical X with a continuous Y

Summary

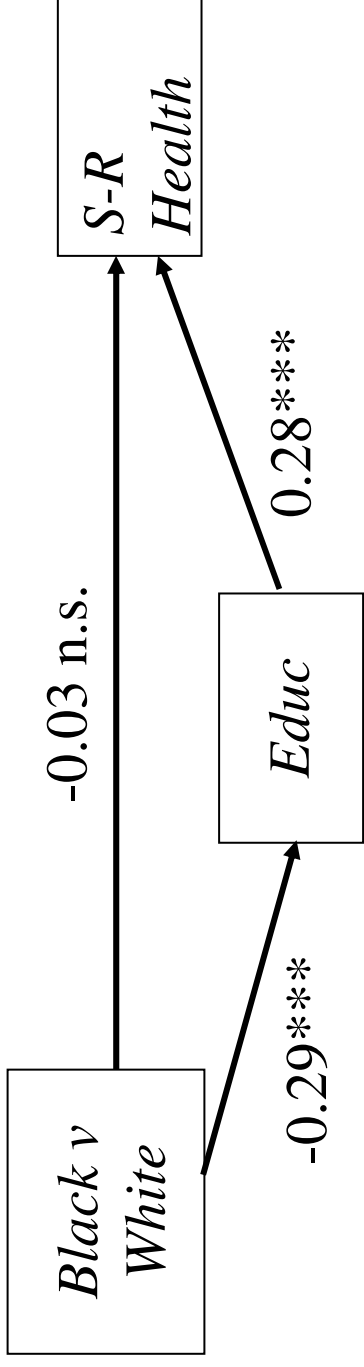
The overall effect was (a) no disparity in self-rated health for among those with less than a HS education, (b) a disparity—favoring Whites, among those with a HS education, and (c) a narrowing of that disparity among those with more than a HS education.

. A significant interaction existed between race and categorical education level. There were no significant differences in self-rated health between Black and White respondents who had less than a HS education. Among those with a HS education, Whites had significantly higher levels of self-rated health, compared to Blacks, $p < .01$. Among those with more than HS education, there was a trend for Whites to have higher self-rated health than Blacks, $p = .071$.

. Mostly, self-rated health significantly increased with each increase in education level. There were two exceptions where an increase in education level was not associated with a significant increase in self-rated health: among Blacks, the increase from 8-11 years to HS education; and among Whites, the increase from HS to >HS education.

Revisiting the mediation model

- The mediation model...



...assumes that the effect of Race is constant at all education levels

. Defending the estimated conditional effect of Race rests upon this assumption

Revisiting the mediation model

- The effect of Race was not constant at all Educ levels
Conversely the effect of Educ is not constant for each Race
- Therefore, the *mediation* model is misspecified, misleading, indefensible.
It estimates the effect of Race conditional on a single effect of Educ
But the effect of Educ is not constant across the races
The mediation model suggests that conditional on Educ,
Race has no direct effect
- The *moderation* model
Showed a significant Race effect at higher levels of Educ
- Interpretation: Race does directly affect General Health,
but only at higher levels of Educ.

There is a price to be paid for ignoring potential interaction effects.

Extensions interaction effects

We have covered interactions between

2 binary variables,

A binary and a continuous variable, and

A binary and a categorical variable

It is also possible to have interactions between

2 categorical variables

2 continuous variables

Aiken, LS & West, SG (1991).

Multiple Regression: Testing and Interpreting Interactions. Sage.

3 or more variables

Parting thoughts

- . Whenever an interaction involves a categorical variable consult the omnibus, multi-df, test of the interaction
 - If significant, explore simple effects within each level of the categorical explanatory variable
 - If there are 2 categorical variables, you have a choice:
 - you can explore the effects of X1 on Y within each category of X2, or
 - you can explore the effects of X2 on Y within each category of X1
 - . Usually, interacting variables are conceptualized as being contemporaneous
 - That is, one variable is not assumed to cause the other.
- Testing interactions may provide important insights into health disparities